Text length and the thematic concentration of text

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ABSTRACT

The impact of text length very often biases results of stylometric indices which are based on rank-frequency distribution (e.g. type-token ratio, repeat rate, entropy). The aim of the article is to observe the relation between text size and thematic concentration indicators \((TC, STC)\). The corpus consists of 1471 English texts of various genres. The obtained results show that thematic concentration is independent of text length in the interval \(<200; 6500>\). Given that the analysis corroborates the findings of the previous research in Czech language, \(TC\) and \(STC\) seem to be reliable stylometric indicators applicable to text analyses of different languages.

Keywords: Normalization, Stylometry, Text analysis, Text length, Thematic concentration

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INTRODUCTION

In quantitative text linguistics, the text length is a factor which influences a majority of methods used for text analysis. The most illustrative is a history of measurement of vocabulary richness; this history can be interpreted as a series of attempts to eliminate an undesirable impact of text length. Typically, some normalization in a computation is used (cf. Popescu et al. 2009, 2011; Wimmer et al. 2003; Těšitelová 1992; Guiraud 1954). The main problems of normalization are as follows: a) particular ways of normalization are not linguistically interpretable (for instance, a use of logarithm or square root in a formula), b) in most cases the normalization does not work well – it somewhat decreases the impact of the text length on a given method, however, it still influences the measurement (cf. Čech 2015). Another way is to use only a part of the text (e.g. first 100 or 1000 words). But this approach does not respect the “integrity” of text and, consequently, it is not adequate, at least from some points of view. Recently, a method called the moving average type-token ratio (introduced by Convington and McFall 2010 and further elaborated by Kubát and Milička 2013) seems to bring promising results in the elimination of text length in the measurement of vocabulary richness. Even though this method is theoretically applicable not only to the measurement of type-token ratio, it has not been employed to other methods yet.

Despite the fact that the text length is indeed a very strong factor influencing particular indices (e.g., the type-token ratio, repeat rate, entropy, lambda structure of text), there are some methods which could be immune to it because of their character. For instance, there is no reason to expect that text activity (cf. Zörnig 2015) should decrease/increase with increasing text length. Further, in some cases, it is possible to determine empirically an interval in which indices are independent of the text length and, importantly, to give theoretical reasons for the determination (cf. Čech 2015).

The later approach seems to be acceptable for the measurement of thematic concentration, as was presented for Czech by Čech (2016). Needless to say, it is necessary to analyse more languages and more texts to explore the
relationship between the thematic concentration and text length. Therefore, we analyse English texts in this study. The article is organized as follows; two methods of measurement of thematic concentration are presented in Section 2, language material used for the analysis is described in Section 3, results are presented in Section 4 and, the article is finalized by Conclusion (Section 4).

METHODS OF MEASUREMENT OF THEMATIC CONCENTRATION

There were introduced several methods for the measurement of thematic concentration in linguistics (Popescu et al. 2009; Čech et al. 2013, 2015). For the purpose of this study, we use the original measurement of thematic concentration (hereinafter TC) (Popescu et al. 2009) and its modification called the secondary thematic concentration (hereinafter STC) (Čech et al. 2013). They are defined\(^1\) as

\[
TC = 2 \sum_{r'=1}^{h} \frac{(h - r')f(r')}{h(h - 1)f(1)},
\]

\[
STC = \sum_{r'=1}^{2h} \frac{(2h - r')f(r')}{h(2h - 1)f(1)},
\]

where \(f(1)\) is the highest frequency in the text, \(h\) is the \(h\)-point (see formula 3) and \(T\) is the number of autosemantics with \(r < h\) (if there are more words with the same frequency in the rank-frequency distribution, \(r'\) can also be represented by the average rank). \(h\)-point is defined as

\[
h = \begin{cases} 
  r_i, & \text{if there is } r_i = f(r_i) \\
  \frac{f(r_i)r_{i+1} - f(r_{i+1})r_i}{r_{i+1} - r_i + f(r_i) - f(r_{i+1})}, & \text{if there is } r \neq f(r)
\end{cases}
\]

\(^1\) Here, we present only formulas; for theoretical aspects of the approach, see references.
where \( r_i \) is a rank and \( f(r_i) \) is the respective frequency of this rank; given that \( r_i \) is the highest number for which \( r_i < f(r_i) \) and \( r_{i+1} \) is the lowest number for which \( r_{i+1} > f(r_{i+1}) \). Thus, if no rank is equal to the respective frequency, one computes the lower part of formula (3) consisting of neighboring values.

**LANGUAGE MATERIAL**

For the analysis, 1471 English texts of various genres were used, see Table 1. The length of texts lies in an interval \( N = <51; 360\,000> \) tokens. All texts were processed by the software *QUITA* (Kubát et al. 2014). For a computation, word forms were used as a unit.

<table>
<thead>
<tr>
<th>genre</th>
<th>number of texts</th>
</tr>
</thead>
<tbody>
<tr>
<td>news</td>
<td>150</td>
</tr>
<tr>
<td>US presidential speeches</td>
<td>57</td>
</tr>
<tr>
<td>scientific texts</td>
<td>60</td>
</tr>
<tr>
<td>short stories</td>
<td>103</td>
</tr>
<tr>
<td>chapters of novels</td>
<td>777</td>
</tr>
<tr>
<td>novels</td>
<td>23</td>
</tr>
<tr>
<td>poems</td>
<td>124</td>
</tr>
<tr>
<td>personal letters</td>
<td>177</td>
</tr>
</tbody>
</table>

**RESULTS**

Before presenting the results, let us consider theoretical aspects of the relationship between the indices (\( TC, STC \)) and the text length. As for extremely short texts (e.g. poems, letters, tweets), the author has a great possibility to control the frequency characteristics of used vocabulary. However, minimal changes of frequency (in an extreme cases only one occurrence or its absence) influence the value of the indices fundamentally. For instance, a short letter (it contains \( N = 45 \) tokens) written by John Keats has the maximal value of the \( TC = 1 \); the absence of sole occurrence of word “Dilke” would mean that the \( TC \) of the
letter equals zero, see Table 2. Consequently, it does not seem to
be adequate to use the method for the analysis of very short
texts. A minimum length of text can be derived empirically, as is
presented below.

Table 2. The rank-frequency distribution of word forms in a short
letter (\(N = 45\) tokens) written by John Keats.

<table>
<thead>
<tr>
<th>Rank</th>
<th>token</th>
<th>frequency</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>Dilke</td>
<td>3</td>
</tr>
<tr>
<td>2</td>
<td>of</td>
<td>2</td>
</tr>
<tr>
<td>3</td>
<td>this</td>
<td>2</td>
</tr>
<tr>
<td>4</td>
<td>shall</td>
<td>2</td>
</tr>
<tr>
<td>5</td>
<td>my</td>
<td>2</td>
</tr>
<tr>
<td>6</td>
<td>given</td>
<td>1</td>
</tr>
<tr>
<td>...</td>
<td>...</td>
<td>...</td>
</tr>
<tr>
<td>45</td>
<td>send</td>
<td>1</td>
</tr>
</tbody>
</table>

The increasing length of text leads to the increasing
frequency of synsemantics (caused by the grammar) and to the
increasing value of the normalizing constant (the divisor in the
formula 1 and 2) which would cause the decrease of the \(TC\) (or
\(STC\)). On the other hand, the increasing length of text leads to
increasing value of the \(h\)-point, consequently, the probability of
occurrence of autosemantics above the \(h\)-point increases, too,
and, consequently, it would cause the increase of the \(TC\) (or
\(STC\)). Thus, there seem to be two opposite mechanisms whose
interaction results in the elimination of the impact of text length
on the thematic concentration.

As for extremely long texts (e.g, novels consisting of
several volumes), there is a minimal chance that the author can
control the frequency characteristics of words. The proportion of
words is probably a result of some background mechanism, cf.
“The longer the text, the more the writer loses his subconscious
control over some proportions and keeps only the conscious
control over contents, grammar, his aim, etc. But as soon as
parts of control disappear, the text develops its own dynamics
and begins to abide by some laws which are not known to the
writer but work steadily in the background. The process is
analogous to that in physics: if we walk, we consider our
activity as something normal; but if we stumble, i.e. lose the control, gravitation manifests its presence and we fall. That means, gravitation does not work ad hoc in order to worry us maliciously, but it is always present, even if we do not realize it consciously. In writing, laws are present, too, and they work at a level which is only partially accessible. One can overcome their working, but one cannot eliminate them. On the other hand, if the writer slowly loses his control of frequency structuring, a new order begins to arise by self-organization or by some not perceivable background mechanism“ (Popescu et al. 2012, 126–127).

The results (see Figure 2–4) confirm the theoretical assumptions mentioned above: the shortest texts have either zero values of indices, or extreme ones. One the other hand, the $TC$ and $STC$ of very long texts lies in a minimal interval.

For Czech, Čech (2016) derived empirically the interval in which the $TC$ and $STC$ are independent on the text length; specifically, the interval is $N = <200; 6500>$ tokens. Analogously, we applied the same procedure and the results are presented in Figures 3 and 4. The linear regression line is almost horizontal, consequently, we can state that the $TC$ and $STC$ are independent on text length in the interval. Obviously, the boundaries of the interval shall be taken as fuzzy ones.

Figure 1. The relationship between the $TC$ and text length in 1471 English texts. The $x$-axis is logarithmic.
Figure 2. The relationship between the STC and text length in 1471 English texts. The x-axis is logarithmic.

Figure 3. The relationship between the TC and text length in 1471 English texts with the length $N = <200; 6500>$ tokens.
Figure 4. The relationship between the \textit{STC} and text length in 1471 English texts with the length $N = \langle 200; 6500 \rangle$ tokens.

CONCLUSION

The study revealed that the values of both the \textit{TC} and \textit{STC} are independent on text length in the interval $N = \langle 200; 6500 \rangle$. Thus, the indices can be “safely” used in this interval for any text analysis of English and Czech texts (cf. Čech 2016 for Czech). We assume that the same results could be obtained for other languages as well because there is no reason to expect some language specific characteristics, i.e., different boundary conditions. Needless to say, this prediction must be corroborated (or rejected) empirically.

REFERENCES


